



## 2 How unusual is the recent series of warm years?

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4 Received 5 October 2008; accepted 18 November 2008; published XX Month 2008.

6 [1] Previous statistical detection methods based partially  
7 on climate model simulations indicate that, globally, the  
8 observed warming lies very probably outside the natural  
9 variations. We use a more simple approach to assess recent  
10 warming at different spatial scales without making explicit  
11 use of climate simulations. It considers the likelihood that  
12 the observed recent clustering of warm record-breaking  
13 mean temperatures at global, regional and local scales may  
14 occur by chance in a stationary climate. Under two statistical  
15 null-hypotheses, autoregressive and *long-memory*, this  
16 probability turns to be very low: for the global records  
17 lower than  $p = 0.001$ , and even lower for some regional  
18 records. The picture for the individual long station records  
19 is not as clear, as the number of recent record years is  
20 not as large as for the spatially averaged temperatures.

21 **Citation:** Zorita, E., T. F. Stocker, and H. von Storch (2008), How  
22 unusual is the recent series of warm years?, *Geophys. Res. Lett.*, 35,  
23 LXXXXX, doi:10.1029/2008GL036228.

### 25 1. Introduction

26 [2] Global mean surface air temperature shows a positive  
27 trend in the 20th century [Brohan *et al.*, 2006; Hansen *et al.*,  
28 2006; Smith and Reynolds, 2005]. After an initial “flat”  
29 development at a pre-industrial level, a clear warming took  
30 place from the 1910s to the 1940s, leveling off in the 1970s.  
31 Since mid 1970s temperatures show a positive trend of  
32 about 0.18°C per decade. This temperature increase and its  
33 geographical patterns have been subject to detection studies  
34 [Hegerl *et al.*, 1996; Barnett *et al.*, 2005], which are  
35 partially based on sophisticated statistical analysis of obser-  
36 vations and climate simulations. Here, we pursue a simpler  
37 and intuitive idea, more accessible to the non-expert, to  
38 estimate the likelihood that the recent observed warming is  
39 consistent with the natural variability.

40 [3] The increasing anthropogenic greenhouse forcing  
41 would, according to simple physical reasoning [Arrhenius,  
42 1896], cause a clustering of record warm years at the end of  
43 the observed record, and in fact, the 13 warmest years in  
44 1880–2006 have all occurred in or after 1990. A clustering  
45 of record years at the beginning of the record would be in  
46 contradiction to anthropogenic greenhouse forcing. The  
47 probability of this clustering occurring by chance at the  
48 end of the record can be estimated under different null  
49 hypotheses of the statistical characteristics of natural vari-  
50 ability. The most simple of those is that the annual values of  
51 mean global temperatures are independent of one another. In

this case, the probability  $p$  of the event  $E$  of finding at least 52  
13 of the largest values of a sequence of 127 independent 53  
random numbers on the last 17 places (year 1990 to 2006) is 54  
 $p(E) = (114!17!)/(127!4!) = 1.25 \times 10^{-14}$ . Such clustering 55  
appears as an extremely improbable random event in a sta- 56  
tionary climate. However, the annual surface air-temperatures 57  
displays a serial correlation, even in a stationary natural 58  
climate, due to processes occurring on the land surface, 59  
ocean, and cryosphere. We conceptualize this natural mem- 60  
ory with two statistical models, namely “short term” and 61  
“long term” memory. Within the former, annual mean 62  
temperature is assumed to be the result of an autoregressive 63  
process, which displays an exponentially decaying auto- 64  
covariance function. In the latter, it is described by *long-* 65  
*term persistence* process with a power-law decay of the 66  
auto-covariance function. Some processes that may be 67  
relevant in the climate context have been shown to display 68  
such type of behavior [Bunde *et al.*, 2005; Rybski *et al.*, 69  
2006]. Although it is quite difficult to ascertain whether a 70  
short time series such as the observed global mean annual 71  
temperature obeys this type of behavior, the present ap- 72  
proach can be readily applied to other more complex null 73  
models. It also offers the advantage that it can be also 74  
extended to those spatial scales using some of the longest 75  
individual station records. 76

### 2. Data and Methods 77

[4] We have analyzed global, regional and long station 78  
temperature records: global annual mean temperature in the 79  
period 1880–2006 from three global temperature data sets - 80  
Hadley Centre-CRU [Brohan *et al.*, 2006], NASA GISS 81  
[Hansen *et al.*, 2006] and NOAA NCDC [Smith and 82  
Reynolds, 2005]; the regional annual temperature means 83  
in the period 1850–2006 from a spatial average of those 84  
grid-cells of the Hadley-Centre-CRU data in 26 geograph- 85  
ical regions [Giorgi and Bi, 2005]; and eleven long station 86  
records constructed by blending original time series kindly 87  
provided by Jones and Moberg [2003], and which end in 88  
year 2000, with the station data from year 2001 onwards 89  
provided by NASA-GISS. Details are given in the auxiliary 90  
material.<sup>1</sup> 91

[5] The parameters of the autoregressive models were 92  
estimated from the observed records in the period up to 93  
1960 to limit the influence of the anthropogenic forcing. For 94  
all but two regional records (South Australia and Southern 95  
South America) an autoregressive model of order one (AR-1) 96  
would be adequate. For these two regional records, the 97  
Durbin-Watson test indicated the presence of autocorrelated 98  
residuals in a AR-1 processes. However, this may be due 99  
to chance, as they appear in 2 out of 26 regional records. 100  
Estimation of confidence intervals for lag-1 autocorrelation 101

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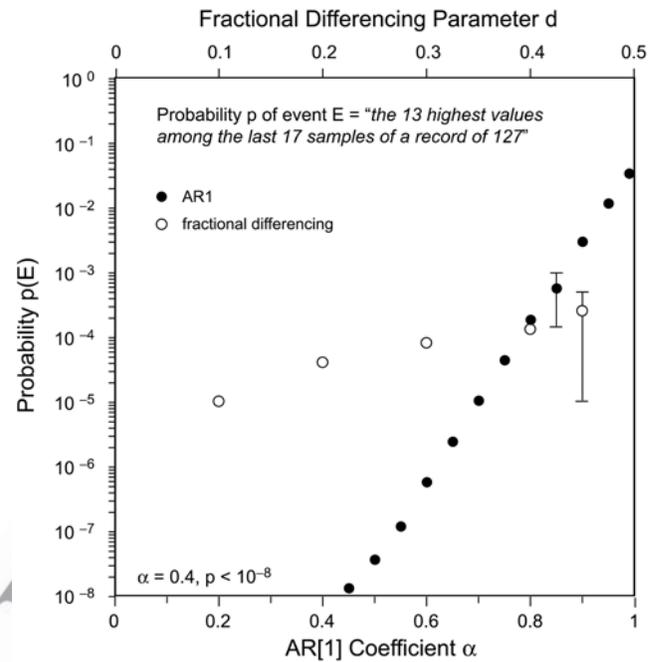
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102 is based on bootstrap methods [Effron and Tibshirani,  
103 1993]; details are given in the auxiliary material.

104 [6] The second null-hypothesis is that the temperature  
105 records are realizations of a ‘long-range autocorrelation’  
106 process [Cohn and Lins, 2005]. The power-law decay of the  
107 autocorrelation is characterized by  $C(k) \sim k^{-\gamma}$  where  $k$  is the  
108 time lag. The value of  $\gamma$  is related to the fractional  
109 differencing parameter  $d$  by  $\gamma = 1 - 2d$ . For the process  
110 to be stationary  $d$  must lie between 0 and 0.5. The Whittle  
111 method [Shimotsu and Phillips, 2005] was used here to  
112 estimate its value. Different statistical test of the stationarity  
113 of the global mean temperature have yielded conflicting  
114 results [Stern and Kauffman, 2000]. The Whittle method  
115 gives values slightly larger than 0.5, even disregarding  
116 the period from 1960 onwards, for all three global records.  
117 Proxy-based reconstructions of the Northern Hemisphere  
118 annual temperature in the past millennium, with varia-  
119 tions less affected by anthropogenic greenhouse forcing than  
120 in the 20th century, yield values for  $d$  in the range 0.32 to  
121 0.54 [Rybski et al., 2006]. Considering these possible  
122 uncertainties, it will be assumed here that  $d$  is smaller, but  
123 very close, to 0.5. It is noted that other more complex null  
124 models are possible, e.g., where both types of short-term  
125 and long-term persistence are present [Cohn and Lins, 2005].  
126 The simultaneous estimation of both parameters becomes,  
127 however, is more problematic. This will be pursued in fur-  
128 ther analysis.

### 129 3. Global Mean Temperature

130 [7] We focus on the number of warm record years in the  
131 last 17 years of the record, as illustrated by Meehl et al.  
132 [2007]. However, this choice is not critical, as the conclu-  
133 sions remain robust when considering the last 10 to last  
134 20 years of the record. Figure 1 shows the probability  $p$ , as  
135 a function of the lag-1 autocorrelation  $\alpha$  or of the fractional  
136 differencing parameter  $d$ , that the clustering of high values  
137 fulfills the criterion that at least  $m$  of the largest values  
138 occupy the last 17 places.  $\log(p)$  under the autoregressive  
139 null-hypothesis exhibits a remarkably linear dependence on  
140  $\alpha$  up to very high values. It is not straightforward to  
141 determine which value of  $\alpha$  would better represent the  
142 natural persistence of the global temperatures, i.e., without  
143 the effect of anthropogenic forcing. The last decades of the  
144 20st century are probably too strongly affected. Tempera-  
145 ture records in the decade 1940–1950 have been found to  
146 be distorted by changing in the measuring devices of sea-  
147 surface temperatures [Thompson et al., 2008] and temper-  
148 ature data in the late 19th century are burdened by higher  
149 uncertainties [Brohan et al., 2006]. To be conservative, i.e.,  
150 risking an overestimation, a value of  $\alpha$  in the range of 0.75  
151 to 0.85 (auxiliary material) can be assumed. Within this  
152 range the probability of a random occurrence of event  $E$   
153 is about  $10^{-5}$  to  $10^{-3}$ , a higher likelihood than under the  
154 null-hypothesis of white noise, but still an extremely rare  
155 event. The dependence of  $\log(p)$  on the value of the  
156 fractional differencing parameter  $d$  is not so steep in the  
157 range of 0 to 0.45, but  $p(E)$  is still quite low for values of  
158  $d$  close to 0.5, yielding a statistical significance for event  
159  $E$  of about  $10^{-3}$  as well. The results for the last  $n = 17$  years  
160 can be considered typical. Figure 1 includes the range of  
161 probabilities obtained when considering the range from  $n =$



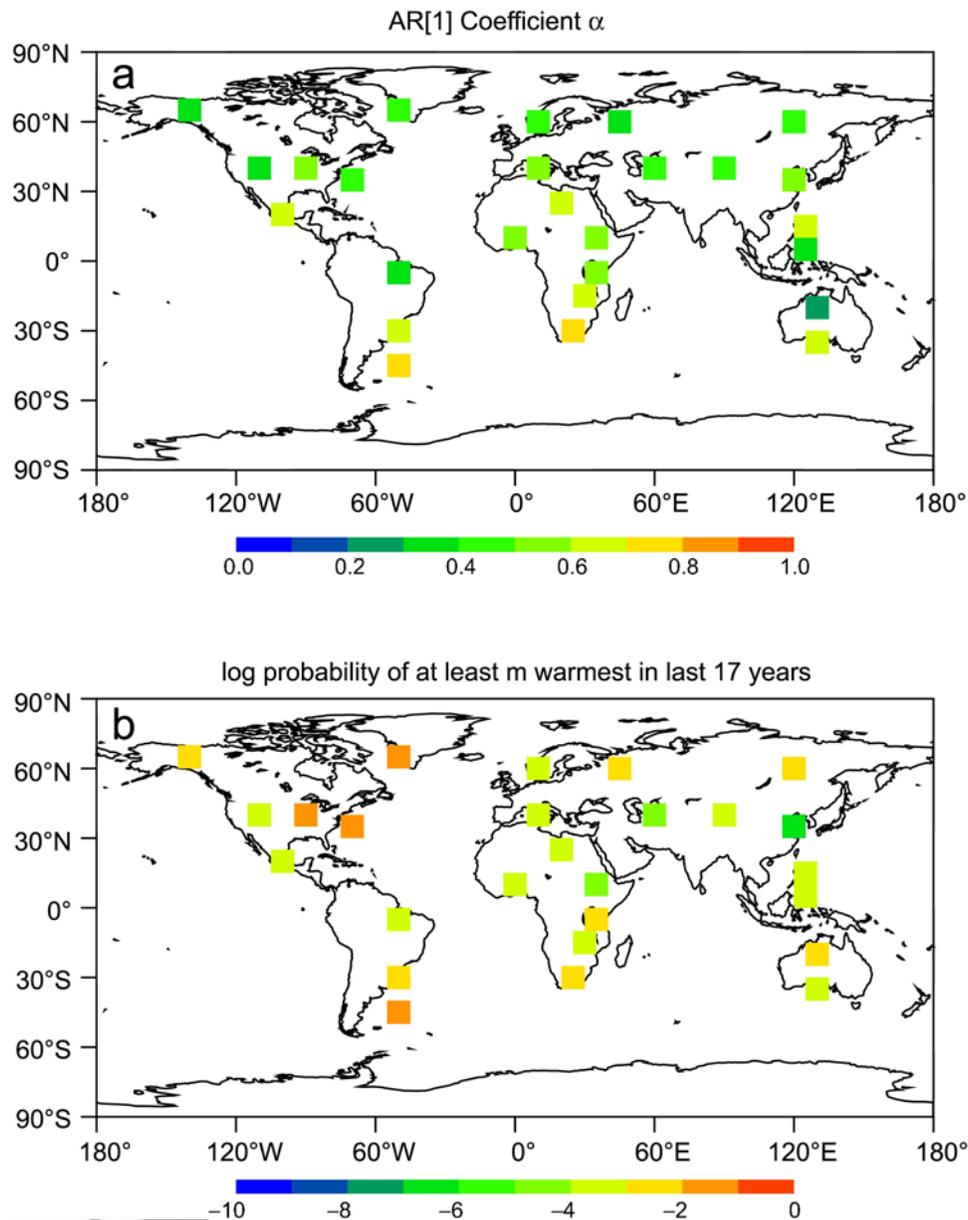
**Figure 1.** Probability that the 13 largest values in synthetic time series of 127 values occupy the last 17 places: closed circles, autoregressive process with different lag-1 autocorrelation  $\alpha$  (bottom axis); open circles, long-range-correlation process for different values of the fractional differencing parameter  $d$  (top axis). Vertical lines around the probability at  $\alpha = 0.85$  and  $d = 0.45$  display the range of probabilities obtained when considering, instead of 17 years, the last 10 to 20 years in the record and, instead of the 13 largest values, their corresponding number of record years as observed in the global mean temperature (see also auxiliary material).

10 to  $n = 20$ , for the cases  $\alpha = 0.85$  and  $d = 0.45$  (auxiliary material). 162 163

### 164 4. Regional Temperature Records

165 [8] Series of record warm years can be also found in  
166 regional temperatures. Here, we consider broad geographi-  
167 cal regions as defined previously [Giorgi and Bi, 2005] and  
168 discussed by Meehl et al. [2007]. The number of warm  
169 record years in the last 17 years is smaller in the regional  
170 series than in the global mean because amplitudes of natural  
171 variability at regional scales are larger. However, the per-  
172 sistence of the regional temperature series is also smaller  
173 (smaller  $\alpha$  and  $d$ ), and therefore the likelihood of the  
174 clustering of a smaller number of record warm years could  
175 still be theoretically smaller than for the global mean tem-  
176 perature, thus making such events even more rare than for  
177 the global mean temperature.

178 [9] Some statistical characteristics of the regional records  
179 are given in the auxiliary material. Figure 2 displays the  
180 value of the lag-1 autoregressive parameter  $\alpha$  and the log-  
181 probabilities of event  $E$  under the autoregressive null-  
182 hypothesis (numerical values are enclosed in the auxiliary  
183 material). The log-probabilities of event  $E$  depend on  $\alpha$ , on  
184 the number of missing values present in the observed



**Figure 2.** (a) Estimated values of the lag-1 autocorrelation parameter  $\alpha$  for the annual mean temperature in each Giorgi region in the period 1850–2007. (b) Log-probability of the event  $E$  that the  $m$  largest values of 157 values occupy the last 17 places in lag-1 autoregressive synthetic time series. For the observed values of  $m$  see the auxiliary material.

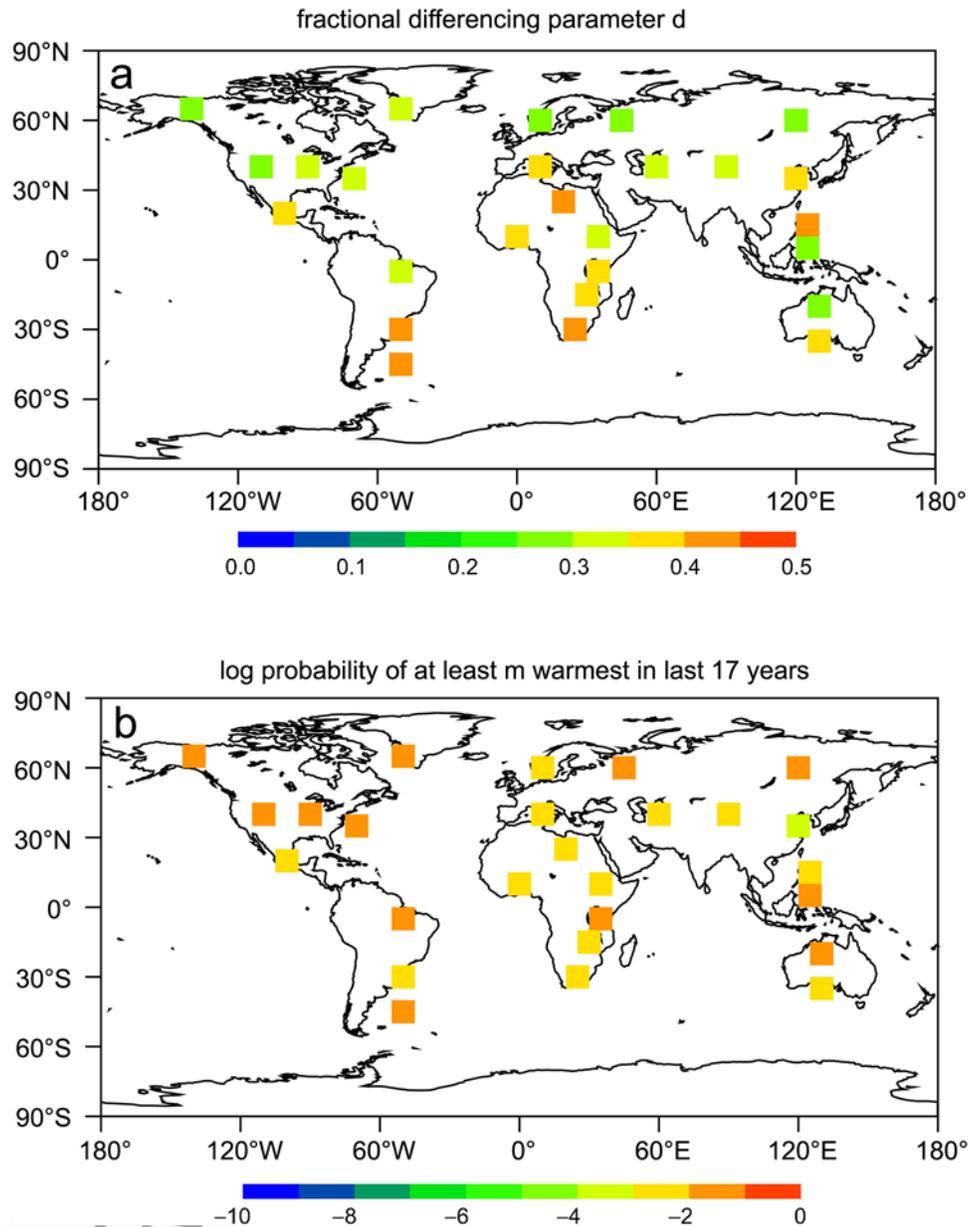
185 records and on the number of warm record years observed  
 186 in the last 17 years. The value of  $\alpha$  shows a slight but  
 187 perceptible meridional gradient with lower values at high  
 188 latitudes in the Northern hemisphere and higher values in  
 189 the regions adjacent to the Southern Ocean. The probabil-  
 190 ities of event  $E$  occurring by chance under the autoregres-  
 191 sive null-hypothesis, on the other hand, do not show a clear  
 192 geographical pattern. For many of the 26 regions the like-  
 193 lihood of event  $E$  under the autoregressive null-hypothesis  
 194 is quite low, especially in Western Europe, Africa and Central  
 195 Asia, in the range  $10^{-5}$  to  $10^{-4}$ , which is even lower than for  
 196 the global mean temperature.

197 [10] The estimated fractional differencing parameter  $d$  for  
 198 the regional records also show a clear geographical pattern,

with a tendency for higher values in the Southern Hemi- 199  
 sphere (Figure 3). The interaction of this gradient with the 200  
 geographical variation of the number of record-years in 201  
 the last 17 years yields no clear pattern. The probability of 202  
 event  $E$  occurring by chance under the long-memory null- 203  
 hypothesis is in general higher than in the autoregressive 204  
 case. Nevertheless, in several regions in Africa and Eurasia, 205  
 this level of significance lies between  $10^{-2}$  and  $10^{-3}$ . 206

## 5. Individual Long Station Records 207

[11] The individual stations are geographically clustered 208  
 in Europe. The results are summarized in Figure 4 and in the 209  
 auxiliary material. Due to the local temperature noise the 210



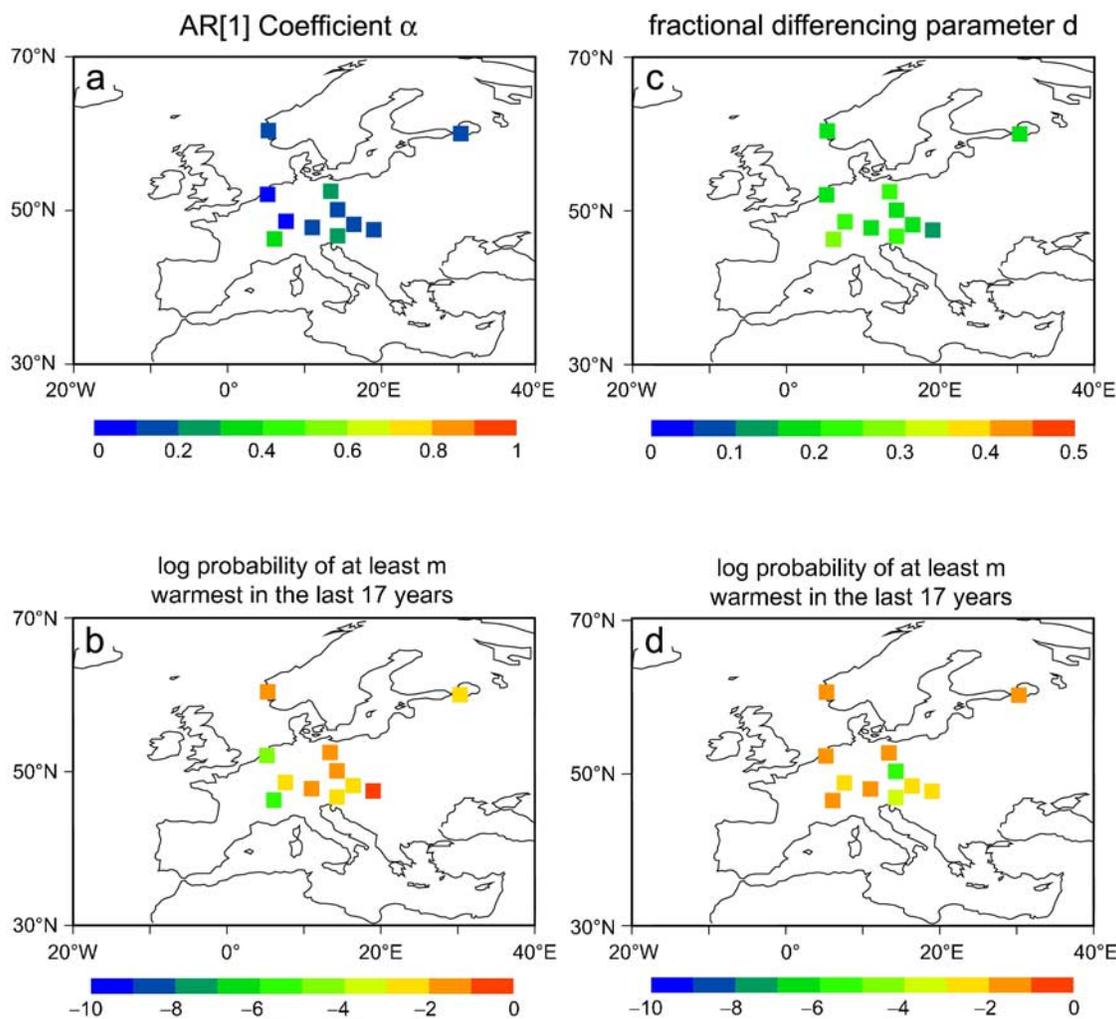
**Figure 3.** (a) Estimated values of the fractional differencing parameter  $d$  for the annual mean temperature in each Giorgi region in the period 1850–2007. (b), Log-probability of the event E that the  $m$  largest values of 157 values occupy the last 17 places in long-term autocorrelation synthetic series. For the observed values of  $m$  see auxiliary material.

211 values of  $\alpha$  are lower than for the regional or global mean  
 212 but the number of recent record years is also not as large.  
 213 Both factors tend to balance, yielding probabilities in the  
 214 same range as for the regional records (Figure 4b). In gen-  
 215 eral, the value of  $d$  for the stations and regional records lies  
 216 within the range 0 to 0.3 (Figure 4c), in accordance to  
 217 independent estimations [Bunde et al., 2005]. The results for  
 218 the long-term-persistence null-hypothesis are qualitatively  
 219 similar to the autoregressive case (Figure 4d).

## 220 6. Conclusions

221 [12] Our assessment of the rarity of the event E depends  
 222 on the underlying null hypothesis. The risk of erroneously  
 223 rejecting the null hypothesis strongly depends on the

assumption about the character of the memory, as is dem- 224  
 onstrated for long-term trends by Cohn and Lins [2005] and 225  
 for the occurrence of extremes by Bunde et al. [2005]. This 226  
 significance can change by orders of magnitude. But, never- 227  
 theless, the clustering of warm years at the end of the 228  
 observed record would appear to be as an extremely rare 229  
 event under either of the two assumptions about the mem- 230  
 ory that considered here. For the global mean temperature, 231  
 conservative values of the lag-1 autocorrelation for global 232  
 temperature yield very low likelihoods for the recent 233  
 clustering of record warmth and under the long-term-persistence 234  
 model this probability is even lower. The analysis of the 235  
 regional records indicates that for some this likelihood is 236  
 lower than for the global mean. This may seem surprising, 237  
 as global records should inherently contain less internal 238



**Figure 4.** (a and b) Same as Figure 2 but for long individual stations records. (c and d) Same as Figure 3 but for long individual stations records.

239 noise and thus a higher signal-to-noise ratio. Two factors are  
 240 regionally at play: a smaller number of record years and a  
 241 smaller statistical persistence, rendering lower likelihoods  
 242 of random clustering. This conclusion also holds for the  
 243 individual station records. For these, the number of recent  
 244 record years and the persistence of the record tend to be  
 245 smaller than for the spatially aggregated records. It should  
 246 be noted that the estimation of  $\alpha$  and  $d$  from the observed  
 247 records could be arguably biased towards larger values (e.g.,  
 248 larger persistence) by the effects of anthropogenic green-  
 249 house forcing, and therefore the estimated probabilities of  
 250 occurrence of event E, under these two null hypotheses,  
 251 could be also biased high.

252 [13] **Acknowledgments.** We thank Armin Bunde for his kind help  
 253 with the long-term-persistence algorithms and A. Moberg and P. Jones for  
 254 permission to use their data products.

## 255 References

256 Arrhenius, S. (1896), On the influence of carbonic acid in the air upon the  
 257 temperature on the ground, *Philos. Mag.*, *41*, 237–276.  
 258 Barnett, T., et al. (2005), Detecting and attributing external influences on  
 259 the climate system: A review of recent advances, *J. Clim.*, *18*, 1291–  
 260 1314.

Brohan, P., J. J. Kennedy, I. Harris, S. F. B. Tett, and P. D. Jones (2006), 261  
 Uncertainty estimates in regional and global observed temperature 262  
 changes: A new data set from 1850, *J. Geophys. Res.*, *111*, D12106, 263  
 doi:10.1029/2005JD006548. 264  
 Bunde, A., J. F. Eichner, J. W. Kantelhardt, and S. Havlin (2005), Long- 265  
 term memory: A natural mechanism for the clustering of extreme events 266  
 and anomalous residual times in climate records, *Phys. Rev. Lett.*, *94*, 267  
 048701, doi:10.1103/PhysRevLett.94.048701. 268  
 Cohn, T. A., and H. F. Lins (2005), Nature's style: Naturally trendy, *Geo-* 269  
*phys. Res. Lett.*, *32*, L23402, doi:10.1029/2005GL024476. 270  
 Effron, B., and R. J. Tibshirani (1993), *An Introduction to the Bootstrap*, 271  
*Monogr. Stat. Appl. Probab.*, vol. 57, Chapman and Hall, New York. 272  
 Giorgi, F., and X. Bi (2005), Updated regional precipitation and temperature 273  
 changes for the 21st century from ensembles of recent AOGCM simula- 274  
 tions, *Geophys. Res. Lett.*, *32*, L21715, doi:10.1029/2005GL024288. 275  
 Hansen, J., M. Sato, R. Ruedy, K. Lo, D. W. Lea, and M. Medina-Elizalde 276  
 (2006), Global temperature change, *Proc. Natl. Acad. Sci. U. S. A.*, *103*, 277  
 14,288–14,293. 278  
 Hegerl, G. C., H. von Storch, K. Hasselmann, B. D. Santer, U. Cubasch, 279  
 and P. D. Jones (1996), Detecting greenhouse-gas-induced climate 280  
 change with an optimal fingerprint method, *J. Clim.*, *9*, 2281–2306. 281  
 Jones, P. D., and A. Moberg (2003), Hemispheric and large-scale surface air 282  
 temperature variations: An extensive revision and update to 2001, *J.* 283  
*Clim.*, *16*, 206–223. 284  
 Meehl, G. A., et al. (2007), Global climate projections, in *Climate Change* 285  
*2007: The Physical Science Basis. Contribution of Working Group I to* 286  
*the Fourth Assessment Report of the Intergovernmental Panel on Climate* 287  
*Change*, edited by S. Solomon et al., pp. 747–845, Cambridge Univ. 288  
 Press, Cambridge, U. K. 289

- 290 Rybski, D., A. Bunde, S. Havlin, and H. von Storch (2006), Long-term  
291 persistence in climate and the detection problem, *Geophys. Res. Lett.*, *33*,  
292 L06718, doi:10.1029/2005GL025591.
- 293 Shimotsu, K., and P. Phillips (2005), Exact local Whittle estimation of  
294 fractional integration, *Ann. Stat.*, *33*, 1890–1933.
- 295 Smith, T. M., and R. W. Reynolds (2005), A global merged land air and sea  
296 surface temperature reconstruction based on historical observations  
297 (1880–1997), *J. Clim.*, *18*, 2021–2036.
- 298 Stern, D. I., and R. K. Kauffman (2000), Detecting a global warming signal  
299 in global temperature series: A structural time series analysis, *Clim.*  
300 *Change*, *47*, 411–438.
- Thompson, D. W. J., K. K. Kennedy, J. M. Wallace, and P. D. Jones (2008), 301  
A large-discontinuity in the mid-twentieth century in observed global- 302  
mean surface temperature, *Nature*, *453*, 646–650. 303
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